

Lecture Topics

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Maximum Likelihood

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For this course you just need to be able to understand broadly the idea and be able to interpret the estimated coefficients and tests

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We move on to look at modelling the conditional variance of returns $Var(R_t|I_{t-1})$

Apple Share Price Return

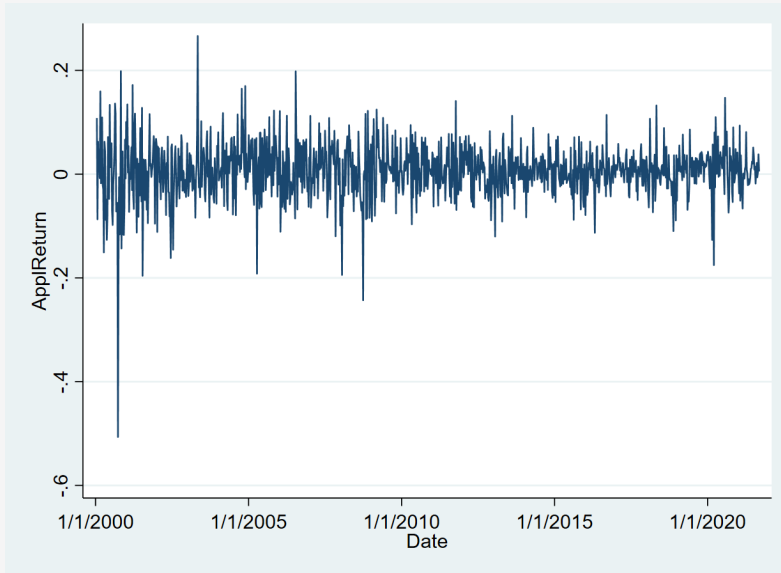


Figure 1: Weekly Apple Stock Return 2000-2021

Apple Share Price Return Correlogram

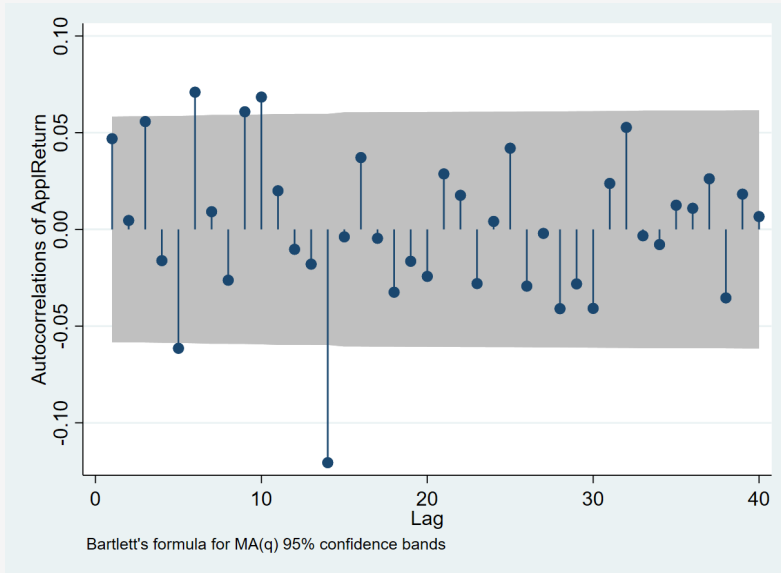


Figure 2: Weekly Apple Price Return Correlogram

Apple Share Price Return Correlogram

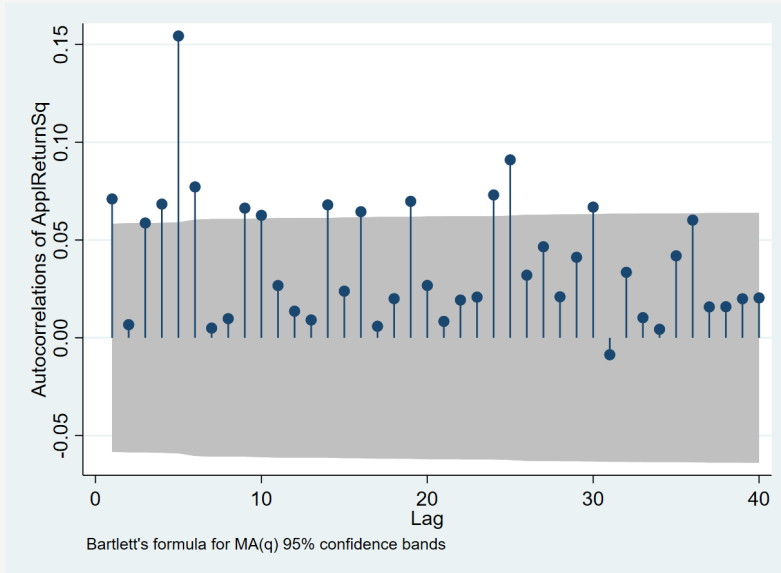


Figure 3: Weekly Squared Apple Price Return Correlogram

Maximum Likelihood

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ML is a general method to estimate the distribution of some variable, which involves estimating a set of parameters.

Maximum Likelihood (ML)

ML assumes we know the function $f(\cdot, \beta)$ is known

Once we know β_0 we know the distribution of y_i

Need to estimate β_0 .

Need to construct an estimator that has "nice" properties like the OLS estimator of a linear model (under the linear model assumptions)

ML estimator achieves this (under assumptions)

Maximum Likelihood (ML)

OLS chooses β to minimise the sum of squared residuals

Many estimators in Econometrics can be expressed as the minimiser/maximiser of some function

For ML want to choose β such that $f(y_i, \beta)$ is as near as possible to the unknown $g(y_i)$

$$K(\beta) = E \left[\log \left(\frac{g(y_i)}{f(y_i, \beta)} \right) \right]$$

Known as Kullback Leibler Info. Criterion

Can show that $K(\beta) \geq 0$ and $K(\beta_0) = 0$

Maximum Likelihood (ML)

If there is a unique β_0 such that $f(y_i, \beta) = g(y_i)$

Then $K(\beta) = 0$ only at β_0

Hence β_0 is the minimiser of $K(\beta)$

For OLS β_0 minimises $E[(y_i - x_i'\beta)^2]$

OLS estimator solves sample version of this

ML minimises sample version of $K(\theta)$

Maximum Likelihood (ML) Idea

$$K(\beta) = E[\log(g(y_i))] - E[\log(f(y_i, \beta))]$$

Hence minimiser of $K(\theta)$ is the minimiser of $-E[\log(f(y_i, \beta))]$

Minimiser of $-E[\log(f(y_i, \beta))]$ is maximiser of $E[\log(f(y_i, \beta))]$

$E[\log(f(y_i, \beta))]$ known as Likelihood Function

Hence the name- Maximum Likelihood Estimation

Maximum Likelihood (ML) Idea

Need to estimate $E[\log(f(y_i, \beta))]$

Use $\frac{1}{N} \sum_{i=1}^N \log f(y_i, \beta)$

$f(y_1, \dots, y_n)$ is the joint likelihood function

$\frac{1}{N} \sum_{i=1}^N \log f(y_i, \beta) = \frac{1}{N} \log(f(y_1, \dots, y_n))$ when y_1, \dots, y_n are i.i.d.

Intuition often given is that β select to maximise probability of observing the sample (y_1, \dots, y_n)

Parameter Inference and Testing

- ↪ To perform inference on any of the estimated parameters we need a distribution.
- ↪ The asymptotic distribution of the $(k \times 1)$ dimensional $\hat{\beta}_{ML}$ is (without any derivation):

$$\hat{\beta}_{ML} \stackrel{a}{\sim} N(\beta_0, \mathcal{I}_n^{-1}(\beta_0))$$

centered around the true but unknown β_0 , where

$$\mathcal{I}_n(\beta_0) = -E\left(\frac{\partial^2 l(\beta_0)}{\partial \beta_0 \partial \beta_0'}\right)$$

is called the information matrix, which is unobserved.

↪ In practice we use the estimated version

$$\hat{\beta}_{ML} \stackrel{a}{\sim} N\left(\beta_0, \mathcal{I}_n^{-1}\left(\hat{\beta}_{ML}\right)\right)$$

where

$$\mathcal{I}_n^{-1}\left(\hat{\beta}_{ML}\right)_{(k \times k)} = - \left(\frac{\partial^2 l\left(\hat{\beta}_{ML}\right)}{\partial \hat{\beta}_{ML} \partial \hat{\beta}'_{ML}} \right)^{-1} \quad (1)$$

and $l\left(\hat{\beta}_{ML}\right)$ is the log-likelihood function evaluated at the ML estimate.

↪ For some $l\left(\hat{\beta}_{ML}\right)$ its derivatives can be written down analytically. Often we require numerical estimates.

- ↔ *Single restrictions*: t - are feasible and asymptotically $N(0, 1)$ distributed. if you use $se_{\hat{\beta}_i}$ from (1)
- ↔ *Multiple restrictions* are tested using a Likelihood-Ratio test

$$LR = 2 \left(\log L \left(\hat{\beta}_u \right) - \log L \left(\hat{\beta}_r \right) \right)$$

where $\hat{\beta}_u$ is the unrestricted parameter estimate and $\hat{\beta}_r$ the parameters estimate with restrictions imposed (e.g. variable exclusions)

- ↔ If restrictions are valid then imposing them should cost very little (in terms of likelihood)
- ↔ Asymptotically $LR \sim \chi_p^2$ where (as usual) p is the number of restrictions imposed

Parameter Estimation

↪ How do we find the ML estimates?

↪ Here only minimal details! The procedure is a sophisticated trial and error strategy which follows the following steps:

1. Propose an initial value for β , $\hat{\beta}_0$ and calculate the log-likelihood function $l(\hat{\beta}_0)$. Declare iteration $0 = best$

2. Iterate over $i = 1, 2, 3, \dots$

2.1 Find an alternative $\hat{\beta}_i = \hat{\beta}_{best} + d$

2.2 calculate the log-likelihood function $l(\hat{\beta}_i)$

2.3 IF $l(\hat{\beta}_i) > l(\hat{\beta}_{best})$ declare $i = best$ and go to next i

ELSE go to 2a and try a few different d

3. Stop the iterations in 2 if no improvements can be found and declare $\hat{\beta}_{ML} = \hat{\beta}_{best}$

↪ The actual implementation of this algorithm differs in how to determine d and when to stop the iterations.